## A Note on Rational Approximation to $(1-x)^{\sigma}$

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THEOREM 1. For  $0 < \sigma \le 1$ ,  $0 \le x \le 1$  and  $n \ge 0$ , we have

$$0 \leq S_n^{-1}(x) - (1-x)^{\sigma} \leq \frac{x^{n+1}}{S_n(x)} \leq {\binom{\sigma+n}{n}}^{-1},$$
(1)

where  $S_n(x) = \sum_{j=0}^n {\binom{\sigma+j-1}{j} x^j}$ . (Note that  $(1-x)^{-\sigma} = \sum_{j=0}^\infty {\binom{\sigma+j-1}{j} x^j}$  for  $0 < \sigma \le 1, \ 0 \le x < 1$ .)

*Proof.* For  $0 < \sigma \le 1$ ,  $0 \le x < 1$ ,  $j \ge 0$ , and  $n \ge 0$ , we obtain, by observing the fact that

$$\binom{n+\sigma+j}{n+1+j} \leq \binom{\sigma+j-1}{j}, \\ 0 \leq S_n^{-1}(x) - (1-x)^{\sigma} = \frac{\sum_{j=n+1}^{\infty} \binom{\sigma+j-1}{j} x^j}{S_n(x)(1-x)^{-\sigma}} \\ = \frac{x^{n+1} \sum_{j=0}^{\infty} \binom{n+\sigma+j}{n+1+j} x^j}{S_n(x)(1-x)^{-\sigma}} \\ \leq \frac{x^{n+1} \sum_{j=0}^{\infty} \binom{\sigma+j-1}{j} x^j}{S_n(x)(1-x)^{-\sigma}} = \frac{x^{n+1}}{S_n(x)}.$$

Let  $0 < \sigma \leq 1$ ,  $n \geq 0$ . As

$$(n+1) S_n(x) > xS'_n(x)$$
 for  $x > 0$ ,  
 $\left(\frac{x^{n+1}}{S_n(x)}\right)' > 0$  for  $x > 0$ ,

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Copyright © 1987 by Academic Press, Inc. All rights of reproduction in any form reserved. and hence  $x^{n+1}/S_n(x)$  is increasing on  $[0, \infty)$ . Thus, for  $0 \le x < 1$ ,

$$0 \leq [S_n(x)]^{-1} - (1-x)^{\sigma} \leq \frac{x^{n+1}}{S_n(x)} < \frac{1}{S_n(1)}$$
$$= \left(\sum_{j=0}^n {\sigma+j-1 \choose j}\right)^{-1} = {\sigma+n \choose n}^{-1}.$$

The last two equalities prove (1) also for x = 1. It follows that

$$\|(1-x)^{\sigma} - S_n^{-1}(x)\|_{L^{\infty}[0,1]} \leq {\sigma+n \choose n}^{-1}.$$
 (2)

*Remarks on Theorem* 1. (i) For  $\sigma = \frac{1}{2}$ , (2) was obtained in [1].

(ii) For  $0 < \sigma < 1$ , (2) is sharper than the estimate given in [2], as for n = 1, 2, ...,

$$\binom{\sigma+n}{n}^{-1} = \frac{\Gamma(\sigma+1)\,\Gamma(n+1)}{\Gamma(\sigma+n+1)} < \frac{n}{n+1}\,\frac{\Gamma(\sigma)\,\Gamma(n)}{\Gamma(\sigma+n)}.$$

THEOREM 2. Let P(x) and Q(x) be polynomials of degree at most  $n \ (\geq 1)$  having only real, non-negative coefficients (Q(0) > 0) and let  $0 < \sigma < 1$ . Then

$$\left\| (1-x)^{\sigma} - \frac{P(x)}{Q(x)} \right\|_{L^{\infty}[0,1]}$$
  
$$\ge (n+1)^{-\sigma} \left( 1 + \left(\frac{n+1}{n}\right)^{n} \right)^{-1} > \frac{(n+1)^{-\sigma}}{1+e}.$$
(3)

Proof. Set

$$\left\| (1-x)^{\sigma} - \frac{P(x)}{Q(x)} \right\|_{L^{\infty}[0,1]} = \varepsilon.$$

Then

$$\varepsilon \ge \left(1 - \frac{n}{n+1}\right)^{\sigma} - \frac{P\left(\frac{n}{n+1}\right)}{Q\left(\frac{n}{n+1}\right)}$$
$$\ge \frac{1}{(n+1)^{\sigma}} + \left(0 - \frac{P(1)}{Q(1)}\right) \left(\frac{n+1}{n}\right)^{n}$$
$$\ge \frac{1}{(n+1)^{\sigma}} - \varepsilon \left(\frac{n+1}{n}\right)^{n}$$

Hence

$$\varepsilon \ge (n+1)^{-\sigma} \left(1 + \left(\frac{n+1}{n}\right)^n\right)^{-1}.$$

*Remarks on Theorem* 2. (i) The estimate (3) is sharper than these obtained in [2, 3]. The technique used here is different from the one used in [3].

(ii) Actually the inequality  $P(1) \ge \max(0, P(n/(n+1)))$  suffices for the proof of Theorem 2 rather than assuming P(x) has only real, non-negative coefficients.

**THEOREM 3.** Let P(x) and Q(x) be polynomials of degree at most  $n (\ge 2)$  and let P(x) have only real, non-negative coefficients. Let  $Q(x) \ne 0$  throughout [0, 1] and let  $0 < \sigma \le 1$ . Then

$$\left\| (1-x)^{\sigma} - \frac{P(x)}{Q(x)} \right\|_{L^{\infty}[0,1]} \ge (16n^{2\sigma})^{-1}.$$

We need the following well known

LEMMA [5, p. 68, Eq. (9)]. If a real polynomial  $Q_n(x)$  of degree  $\leq n$   $(n \geq 0)$  satisfies the inequality  $|Q_n(x)| \leq L$  on [a, b]  $(-\infty < a < b < \infty)$ , then at each real x outside [a, b] we have

$$|Q_n(x)| \leq L \left| T_n\left(\frac{2x-a-b}{b-a}\right) \right|,$$

where  $T_n$  is the nth degree Chebychev polynomial of the first kind.

Proof of Theorem 3. Suppose

$$\left\| (1-x)^{\sigma} - \frac{P(x)}{Q(x)} \right\|_{L^{\infty}[0,1]} < (16n^{2\sigma})^{-1}.$$
(4)

Then on  $[0, 1-n^{-2}]$ ,

$$\left|\frac{P(x)}{Q(x)}\right| > (1-x)^{\sigma} - \frac{1}{16n^{2\sigma}} \ge \frac{1}{n^{2\sigma}} - \frac{1}{16n^{2\sigma}} = \frac{15}{16n^{2\sigma}},$$
$$|Q(x)| < \frac{16n^{2\sigma}}{15} P(1).$$
(5)

In (5) we have used the assumption that P(x) has only real, non-negative coefficients. Now, by applying the lemma to (5) with a = 0,  $b = 1 - n^{-2}$ , we

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get on [0, 1], by familiar elementary properties of  $T_n$ , and by [4, p. 38, Problem 172],

$$\begin{aligned} \max_{\{0,1\}} |Q(x)| &\leq \frac{16n^{2\sigma}}{15} P(1) T_n \left(\frac{n^2 + 1}{n^2 - 1}\right) \\ &< \frac{16n^{2\sigma}}{15} P(1) \left(\frac{n^2 + 1}{n^2 - 1} + \sqrt{\left(\frac{n^2 + 1}{n^2 - 1}\right)^2 - 1}\right)^n \\ &= \frac{16n^{2\sigma}}{15} P(1) \left(\frac{n + 1}{n - 1}\right)^n \\ &\leq \frac{16n^{2\sigma}}{15} P(1) \cdot 9 = \frac{48}{5} n^{2\sigma} P(1). \end{aligned}$$
(6)

From (6) and (4) we obtain  $\frac{5}{48} < n^{2\sigma} |P(1)/Q(1)| < \frac{1}{16}$ , which is false.

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